# Joint Cumulants of *polykays*

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# Introduction

***Abstract*:** Given a multivariate random sample, this set of functions expresses the joint cumulants of multivariate polykays in terms of the joint cumulants of the multivariate population from which the sample is drawn.

Multivariate polykays are minimum-variance unbiased estimators of products of joint cumulants. These estimators are usually indexed by a list of integer vectors, with each vector corresponding to a joint cumulant.

When the index consists of a list of single integers, the multivariate polykay reduces to a univariate polykay. When it consists of a single integer vector, the multivariate polykay reduces to a multivariate *k*-statistic. Finally, when the index consists of a single one-dimensional integer vector, the multivariate *k*-statistic reduces to a univariate *k*-statistic.

***Application Areas/Subject:*** Combinatorics & algebraic methods in statistics.
***Keyword:*** umbral calculus, symmetric polynomial, set partition, multiset, cumulant, *k*-statistic, polykay.
***See Also:*** Maple algorithm [1] and [2]
**Remark:** *k*-statistics, polykays and their multivariate generalization are commonly given in terms of power sums

$$S_{j,\,k,\ldots} = \sum_{i=i}^{n} X_i^j \; Y_i^k \ldots.$$

with $n$ the sample size.

# Initialization

> *restart*

> *with*(*combinat*, *partition*, *multinomial*, *stirling2*)

[ *partition*, *multinomial*, stirling2 ] **(2.1)**

> *echo_type_KS* := *false*;

*echo_type_KS* := *false* **(2.2)**

# Background Functions

The *polyk* function implements four different algorithms. Depending on the input parameters, it returns univariate or multivariate polykays, as well as univariate or multivariate *k*-statistics, expressed in terms of power sums.

Specifically, if the input consists of a single vector containing a single integer, the function returns univariate *k*-statistics. If the input is a single vector containing more than one integer, it returns multivariate *k*-statistics. If the input is a list of vectors, each containing a single integer, it returns univariate polykays. Finally, if the input is a list of vectors, each containing more than one integer, it returns multivariate polykays.

All these functions are available in the Maple worksheet [1] and have been included here for the user's convenience. For the multivariate case, the algorithm that generates all subdivisions of a multiset has also been included. This algorithm is presented and discussed in the Maple worksheet [2]. For further details on the procedures and their mathematical background, see references [3–6].

## *K*-Statistics

The *n*-th *k*-statistic $k_n$ is the unique symmetric unbiased estimator of the *n*-th cumulant $\kappa_n$ of a given statistical distribution, that is

$$E[k_n] = \kappa_n$$

where *E* denotes the expectation. They are expressed in terms of power sums involving the random variables of a random sample.

>

> *makeTab_ss* := **proc**(*N*)
**local** *x*, *y*, *i*;

$$\left[ seq\left( \left[ \frac{N!}{mul\left(x!^{numboccur(y,x)} \cdot numboccur(y,x)!, x = \{op(y)\}\right)} \cdot mul\left(k_i, i = y\right), mul\left(S_i, i = y\right) \right], y = partition(N) \right) \right]$$

**end proc**:

```
makeK_s := proc(N) local i, j;
    [seq(k_i = add(stirling2(i, j)·x^j·(-1)^(j-1)·(j-1)!, j = 1..i), i = 1..N)]
end proc:

fd := proc(j, k) local i; expand(mul(n - i - k, i = 0..j - k)) end proc:
ks := proc(N) local u, v, i; global x;

    v := expand(eval(makeTab_ss(N), makeK_s(N)));
    u := [seq(x^i = (-1)^(i-1)·(i-1)!·fd(N - 1, i), i = 1..N)];
    add(x_1·x_2, x = expand(eval(v, u)))
    ------------------------------------
          mul((n - x), x = 0..N - 1)
end proc:
```

Exampe: The *k*-statistic of order 3 is

```
> ks(3)
```

$$\frac{n^2 S_3 - 3\, n\, S_1 S_2 + 2 S_1^3}{n\,(n-1)\,(n-2)} \qquad \textbf{(3.1.1)}$$

## Polykays

The symmetric statistic $k_{r,s,..}$ is an unbiased estimator of the cumulant product $\kappa_r\,\kappa_s...$, that is

$$E[k_{r,s,..}] = \kappa_r\,\kappa_s...$$

where $E$ denotes the expectation and $\kappa_r$ is the *r*-th cumulant. These statistics are called polykays or generalized *k*-statistics.

```
> makeCTR_s := proc(N)
  local i, v, x, r;
      [seq(k_i = add(mul(μ_r, r = v)·(-1)^(nops(v)-1)·(nops(v) - 1)!i!
      ---------------------------------------------------------------, v = partition(i)), i
      mul((x)!^numboccur(v,x)·numboccur(v, x)!, x = {op(v)})
      = 1..N)]
  end proc:


  makeMu := proc( )
      local u, v, N, eu, i, x, k;
      N := add(i, i = args);
      eu := [seq(μ_i = 1, i = 1..N)];
      if nargs = 1 then
          u := [seq([[x]], x = partition(args_1))]
      else
          u := [seq([seq(x, x = partition(i))], i = args)]
      end if;
```

```
    u := add(mul((-1)^(nops(i)-1)·(nops(i)-1)!·mul(μ_k, k=i)·countP(i), i=v)·fd(N-1,
    add(nops(k), k=v))·countP([seq(op(i), i=v)])^(-1), v=`if`(nargs=1, comb(u, 1, []),
    [comb(u, 1, [])]));
     if nops(indets(u))=1 then u := [u]; fi;
    seq(simplify(v/eval(v, eu)) = eval(v, eu), v=u)
 end proc:


comb :=proc(V, ptr, Y)
    local L;if ptr=nops(V)+1 then
        return Y
    end if;
    seq(comb(V, ptr+1, [op(Y), L]), L=V_ptr)
end proc:

countP := proc(u)
     local x;
                          add(x, x=u)!
     ------------------------------------------------------
     mul(x^(numboccur(u, x))·numboccur(u, x)!, x={op(u)})
end proc:

ps := proc()
 local u, v, N, x;
    N := add(x, x=args);
     if mul(x, x=args)=0 then error "A null parameter is not permitted."; fi;
    u := expand(eval(makeTab_ss(N), makeCTR_s(N)));
    v := expand(eval(u, [makeMu(args), μ=0]));
          add(x_1·x_2, x=v)
    ---------------------------
    mul(n-x, x=0..N-1)

 end proc:
```

Example : The polykay of order (2, 1) is

```
> ps(2, 1);
```

$$\frac{-S_1^3 + (1+n)\,S_1\,S_2 - n\,S_3}{n\,(n-1)\,(n-2)} \qquad \textbf{(3.2.1)}$$

## Multiset subdivision

The *makeTab* function lists all subdivisions of a multiset (for a thorough discussion see [2]). With multivariate *k*-statistics and/or multivariate polykays, this function speeds up the overall procedure (see [7]). For details on the procedure and mathematical background see also [3] and [4].

```
> nRep :=proc(u) local x; mul(x_2!, x = convert(u, multiset)) end proc:
```

```
URv := proc(u, v)
    local U, ou, i, ptr, vI;
    ou := NULL; U := [ ]; vI := indets(v);
    for ptr from nops(u) by -1 to 2 do
        if has(u[ptr], v) then break end if
    end do;
    for i from ptr to nops(u) do
        if not (u[i] = ou or has(u[i], vI)) then
            ou := u[i];
            U := [op(U), [op(u[1..i-1]), u[i]·v, op(u[i+1..-1])]]
        end if
    end do;
    op(U), [op(u), v]
end proc:
```

```
URV := proc( )
    local U, V, i, u, x;
    U := [args[1,1]]; V := args[2,1];
    for i to nops(V) do U := [seq(URv(u, V[i]), u = U)] end do;
    seq([x, args[1,2]·args[2,2]], x = U)
end proc:
```

```
URmV := proc( )
    local U, i, u, x;
    if nargs = 1 then args else
        U := URV(args[1], args[2]);
        for i from 3 to nargs do
            U := seq(URV(u, args[i]), u = [U])
        end do;
        seq([x[1], x[2]/nRep(x[1])], x = U)
    end if
end proc:
```

```
makeTab := proc( )
    local U, x, i, y, z, r;
    if add(x, x = args) = 0 then
        return 0
    end if;
    U := [seq(`if`(args[i] = 0, NULL, [seq([[seq((P || i)^z, z = y)]], multinomial(args[i], seq(r, r
```

```
        =y) ) ], y=partition(args_i) ) ]), i=1..nargs) ];
    if nops(U) = 1 then
        [seq([x_1, x_2/nRep(x_1)], x=op(U))]
    else
        [seq(URmV(op(x)), x=[comb(U, 1, [ ])])]
    end if
end proc:
>
```

## Multivariate *k*-statistics

The multivariate *k*-statistic is the unique symmetric unbiased estimator of the *n*-th joint cumulant $\kappa_{i,j,..}$ of a given multivariate statistical distribution, that is

$$E[k_{i,j,..}] = \kappa_{i,j,..}$$

where *E* denotes the expectation. They are expressed in terms of multivariate power sums involving the random vectors of a random sample.

```
> makeTab_sm := proc( )
    local vP, U, i, y, x;
    U := makeTab(args);
    vP := [seq(P||i, i=1..nargs)];
    [seq([mul(k_add(degree(x, vP_i), i=1..nops(vP)), x=y_1)·y_2, mul(S_seq(degree(x, vP_i), i=1..nops(vP)), x
    =y_1)], y=U)]
end proc:

km := proc( ) local u, i, v, N; global x;

    N := add(x, x=args);
    v := expand(eval(makeTab_sm(args), makeK_s(N) ) );
    u := [seq(x^i = (-1)^(i-1)·(i-1)!·fd(N-1, i), i=1..N)];
    add(x_1·x_2, x=expand(eval(v, u)))/mul(n-x, x=0..N-1)
end proc:
```

Example : The multivariate $k-$statistics of order $(2, 1)$ is

```
> km(2, 1);
```

$$\frac{n^2 S_{2,1} - n S_{2,0} S_{0,1} - 2 n S_{1,1} S_{1,0} + 2 S_{1,0}^2 S_{0,1}}{n (n-1) (n-2)} \qquad \textbf{(3.4.1)}$$

Note that the power sums in (3.4.1) are:

$$S_{2,1} = \sum_{i=i}^{n} X_i^2 Y_i \quad S_{0,1} = \sum_{i=i}^{n} Y_i \quad S_{2,0} = \sum_{i=i}^{n} X_i^2 \quad S_{1,0} = \sum_{i=i}^{n} X_i \quad S_{1,1} = \sum_{i=i}^{n} X_i Y_i$$

where $(X_1,..., X_n)$ refers to the population $X$ and $(Y_1,..., Y_n)$ refers to the population $Y$.

## Multivariate Polykays

The multivariate polykay $k_{r,s,...;i,j,...;...}$ is an unbiased estimator of the cumulant product $\kappa_{r,s,...}\ \kappa_{i,j,...}...$, that is

$$E[k_{r,s,...;i,j,...;...}] = \kappa_{r,s,...}\ \kappa_{i,j,...}...$$

where $E$ denotes the expectation and $\kappa_{r,s,...}\ \kappa_{i,j,...}$ are joint cumulants of order *(r,s,...)* and *(i,j,..)* respectively.

```
> makeTab_mm := proc( )
    local vP, U, y, x, i;
    U := makeTab(args);
    vP := [seq(P||i, i = 1 ..nargs)];
    [seq([mul(k[seq(degree(x, vP[i]), i = 1 ..nops(vP))], x = y[1])·y[2], mul(S[seq(degree(x, vP[i]), i = 1 ..nops(vP))], x = y[1])], y
      = U)]
  end proc:


  ctr := proc( )
      local vP, U, i, v, x;
      U := makeTab(args);
      vP := [seq(P||i, i = 1 ..nargs)];
      add(v[2]·(-1)^(nops(v[1])-1)·(nops(v[1]) - 1)!·mul(μ[seq(degree(x, vP[i]), i = 1 ..nops(vP))], x = v[1]), v
      = makeTab(args))
  end proc:


  makeCTR_m := proc( )
      local i, y, x; [seq(k[op(i)] = ctr(op(i)), i = comb([seq([seq(x, x = 0 ..y)], y = args)], 1, [ ]))]
  end proc:


  unionVects := proc(U::list, V::list) local v, u;
  if nops(U) = 0 then V
      elif nops(V) = 0 then U
          else [seq(seq([sort([op(v[1]), op(u[1])]), v[2]·u[2]], u = U), v = V)]
```

```
    end if
end proc:
```

```
ricVtab :=proc(v, V, N)
    local u, vv;
    vv := sort(v_1);
    for u in V do
        if sort(u_1) = vv then
            return v_2 · fd(N − 1, nops(v_1)) / u_2
        end if
    end do;
    return 0
end proc:
```

```
pm :=proc( ) local N, M, u, v, i, vK, vTab, vP, vParts, vEval, k, h, y; global x;
    M := max(seq(nops(x), x = args));
    vP := [seq(P||i, i = 1 ..M)];
    u := op(add(k, k = seq([op(h), seq(0, y = nops(h) ..M − 1)], h = args)));
    v := expand(eval(makeTab_mm(u), makeCTR_m(u)));
    N := add(h, h = u);
    vTab := [ ];
    for i to nargs do
        vTab := unionVects(vTab, [seq([v_1, v_2 · (−1)^(nops(v_1)−1) · (nops(v_1) − 1)!], v
        = makeTab(op(args_i)))])
    end do;
    if nops(vTab) > 1 then
        vTab := [seq([op(i)_1, op(i)_2 / 2], i = mul(u_1^(2·u_2), u = vTab))]
    end if;
    vParts := makeTab(u);
    vTab := [seq([x_1, ricVtab(x, vParts, N)], x = vTab)];
    vEval := [seq(mul(μ_(seq(degree(x, vP_i), i = 1 ..nops(vP))), x = y_1) = y_2, y = vTab), μ = 0];
    add(x_1 · x_2, x = eval(v, vEval)) / mul(n − x, x = 0 ..N − 1)
end proc:
```

Example : The multivariate polykay of order $([1, 1], [1, 0])$ is

```
> pm([1, 1], [1, 0]) ;
```

$$\frac{n\,S_{1,1}\,S_{1,0}-S_{1,0}^{2}\,S_{0,1}-n\,S_{2,1}+S_{2,0}\,S_{0,1}}{n\,(n-1)\,(n-2)} \tag{3.5.1}$$

Note that the vectors in the input list must be the same length: if any vector is "smaller," the function adds 0 until it reaches the correct length.

## Master function "polyk" handling all cases

This function allows us to recall all functions for generate *k*-statistics, polykays and their multivariate generalizzations.
The input is the following:

- for generate *k*-statistics $k_r$ the parameter is: [ r ]
- for generate polykays $k_{r,s}$ the parameter is: [ r ], [ s ]
- for generate multivariate *k*-statistics $k_{r,\,s}$ the parameter is: [ r, s ]
- for generate multivariate polykays $k_{r,s;\,u,v}$ the parameter is: [ r, s ], [ u, v]

```
> polyk:=proc( )
    global x, echo_type_KS;
    if echo_type_KS then
      if nargs = 1 then if nops(args[1]) = 1
                        then print(execute_ks(op(args[1])))
                        else print(execute_km(op(args[1]))) end if;
     elif add(`if`(nops(x) = 1, 0, 1), x = args) = 0 then
          print(execute_ps(seq(op(x), x = args)))
     else print(execute_pm(args) )
      end if:
    end if;
   if nargs = 1 then
       if nops(args[1]) = 1 then ks(op(args[1]))
                            else km(op(args[1]))
       end if;
    elif add(`if`(nops(x) = 1, 0, 1), x = args) = 0 then
         ps(seq(op(x), x = args))
    else pm(args)
    end if
  end proc:
```

Esample: The *k*-statistic of order 3 is

```
> polyk([3])
```

$$\frac{n^{2}\,S_{3}-3\,n\,S_{1}\,S_{2}+2\,S_{1}^{3}}{n\,(n-1)\,(n-2)} \tag{3.6.1}$$

Example : The polykay of order $(2, 1)$ is

```
> polyk([2], [1])
```

$$\frac{-S_{1}^{3}+(1+n)\,S_{1}\,S_{2}-n\,S_{3}}{n\,(n-1)\,(n-2)} \tag{3.6.2}$$

Example : The multivariate $k-$statistics of order $(2, 1)$ is

> *polyk*([2, 1])

$$\frac{n^2 S_{2,1} - n S_{2,0} S_{0,1} - 2 n S_{1,1} S_{1,0} + 2 S_{1,0}^2 S_{0,1}}{n (n-1) (n-2)} \quad \textbf{(3.6.3)}$$

Example : The multivariate polykay of order $([1, 1], [1, 0])$ is

> *polyk*([1, 1], [1])

$$\frac{n S_{1,1} S_{1,0} - S_{1,0}^2 S_{0,1} - n S_{2,1} + S_{2,0} S_{0,1}}{n (n-1) (n-2)} \quad \textbf{(3.6.4)}$$

# Replacing symbols with numerical data

The following functions allow the random variables of a random sample to be replaced with corresponding numerical observations.

The *powS* function returns the sum of the *r*-th powers of the observations.

```
> powS :=proc( )
    local j; if nargs = 1 then
        Sum('X_i^args^1, i = 1 ..'n')
    else
        Sum(mul('X_{i,j}^args^j, j = 1 ..nargs), i = 1 ..'n')
    end if
  end proc:
```

Example

> *powS*(5, 3, 1)

$$\sum_{i=1}^{n} X_{i,1}^5 X_{i,2}^3 X_{i,3} \quad \textbf{(3.6.1.1)}$$

> *powS*(9);

$$\sum_{i=1}^{n} X_i^9 \quad \textbf{(3.6.1.2)}$$

The *npolyk* function computes a numerical value of a *k*-statistic or a polykay (including the multivariate cases), replacing the random variables of the random sample with the corresponding numerical observations. The parameters are the following:

- for generate *k*-statistics $k_r$ the parameters are: [ r ], [ [ n1, n2, ...] ]
- for generate polykays $k_{r;s}$ the parameters are: [ [ r ], [ s ] ], [ [ n1, n2, ...] ]
- for generate multivariate *k*-statistics $k_{r,s}$ the parameters are: [ [ r , s ] ], [ [ n1a, n2a], [ n1b, n2b] , ... ]

- for generate multivariate polykays $k_{r,s;\,u,v}$ the parameters are: [ [ r , s ], [ u , v] ], [ [ n1a, n2a], [ n1b,

n2b] , ... ]

```
> npolyk:=proc(V, data)
    local res, ind, N, vE, Si, y, j, x;
    Si := false;
    if nops(V) = 1 then
        if nops(op(V)) = 1 then
            Si := true;
            N := nops(op(data));
            res := ks(op(op(V)))
        else
            N := nops(data);
            res := km(op(op(V)))
        end if;
    elif add(`if`(nops(x) = 1, 0, 1), x = V) = 0 then
        Si := true;
        N := nops(op(data));
        res := ps(seq(op(x), x = V))
    else
        N := nops(data);
        res := pm(op(V))
    end if;
    ind := `minus`(indets(res), {n});
    vE := seq(y[1] = Sum(mul('X'[`if`(Si, op([j, i]), op([i, j]))]^y[2,j], j = 1 ..nops(y[2])), i = 1 ..N),
    y = seq([x, [op(x)]], x = ind));
    eval(res, [eval(evalf(vE), [X = data]), n = N])
end proc:
```

Examples: suppose to have the following (univariate) numerical sample

```
> data := [[16.34, 10.76, 11.84, 13.55, 15.85, 18.20, 7.51, 10.22, 12.52, 14.68, 16.08, 19.43, 8.12,
    11.20, 12.95, 14.77, 16.83, 19.80, 8.55, 11.58, 12.10, 15.02, 16.83, 16.98, 19.92, 9.47, 11.68,
    13.41, 15.35, 19.11]]
```

$$data := [[16.34, 10.76, 11.84, 13.55, 15.85, 18.20, 7.51, 10.22, 12.52, 14.68, 16.08, 19.43, 8.12, 11.20, 12.95, 14.77, 16.83, 19.80, 8.55, 11.58, 12.10, 15.02, 16.83, 16.98, 19.92, 9.47, 11.68, 13.41, 15.35, 19.11]] \quad \textbf{(3.6.1.3)}$$

An estimation of the mean (*k*-statistic of first order) is

```
> npolyk([[1]], data)
```

$$14.02166667 \quad \textbf{(3.6.1.4)}$$

An estimation of the variance (*k*-statistic of second order) is

```
> npolyk([[2]], data)
```

$$12.65006954 \quad \textbf{(3.6.1.5)}$$

An estimation of the skewness is $k_3 / \sqrt{k_2^3}$

> $\frac{npolyk([[3]], data)}{\sqrt{npolyk([[2]], data)^3}}$

$$-0.03216232203 \qquad \textbf{(3.6.1.6)}$$

An estimation of the kurtosis is $k_4 / k_2^2$

> $\frac{npolyk([[4]], data)}{npolyk([[2]], data)^2}$

$$-0.8852932701 \qquad \textbf{(3.6.1.7)}$$

The estimation of the product $\kappa_3\, \kappa_2$ is

> $npolyk([[3], [2]], data)$

$$-15.56090621 \qquad \textbf{(3.6.1.8)}$$

Examples: suppose to have the following (bivariate) numerical sample

> $data := [[5.31, 11.16], [3.26, 3.26], [2.35, 2.35], [8.32, 14.34], [13.48, 49.45], [6.25, 15.05], [7.01, 7.01], [8.52, 8.52], [0.45, 0.45], [12.08, 12.08], [19.39, 10.42]]$

$$data := [[5.31, 11.16], [3.26, 3.26], [2.35, 2.35], [8.32, 14.34], [13.48, 49.45], [6.25, 15.05], [7.01, 7.01], [8.52, 8.52], [0.45, 0.45], [12.08, 12.08], [19.39, 10.42]] \qquad \textbf{(3.6.1.9)}$$

An estimation of the joint cumulant $\kappa_{2,1}$ is

> $npolyk([[2, 1]], data)$

$$-23.73790303 \qquad \textbf{(3.6.1.10)}$$

An estimation of the product of the joint cumulant $\kappa_{2,1}\ \kappa_{1,0}$ is

> $npolyk([[1, 1], [1, 0]], data)$

$$294.2657621 \qquad \textbf{(3.6.1.11)}$$

An estimation of the product of the joint cumulant $\kappa_{2,1}\ \kappa_{2,0}\, \kappa_{1,0}$ is

> $npolyk([[2, 1], [2, 0], [1, 0]], data)$

$$12369.47541 \qquad \textbf{(3.6.1.12)}$$

# Joint Cumulant of polykays: auxiliary functions

The expressions produced by the following set of functions correspond to formula (4.10) on p. 96 of [8]. This formula expresses the cumulants of multivariate polykays in terms of the joint cumulants of the population underlying the multivariate sample. This representation was chosen because it reduces the overall complexity of the final result.

Considering first the simplest case of *k*-statistics, the procedure is as follows:

1. The cumulants of the *k*-statistics are expressed in terms of their moments using the function *ctr*.
2. The moments of the *k*-statistics are expressed in terms of the moments of the population underlying the sample using the functions *pSS* and *E*.

3. The moments of the population underlying the sample are expressed in terms of its cumulants using the function *rct*.

The same procedure applies to multivariate polykays.

Further details on each step are provided below.

## Cumulants and Moments

The *ctr function* (see [1]) expresses cumulants of a random vector in terms of joint moments $\mu_{i,j,\ldots}$. Details on the procedure and mathematical background are given in [1] and [3]

Example: The joint cumulant $\kappa_{2,1}$ in terms of joint moments is

```
> ctr(2, 1)
```

$$2\,\mu_{1,0}^{2}\,\mu_{0,1} - \mu_{2,0}\,\mu_{0,1} - 2\,\mu_{1,1}\,\mu_{1,0} + \mu_{2,1} \qquad \textbf{(4.1.1)}$$

The *rct* function expresses joint moments of a random vector in terms of joint cumulants $\kappa_{i,j,\ldots}$. Details on the procedure and mathematical background are given in [3].

```
> rtc := proc( )
  local vP, U, i, v, x;
      U := makeTab(args);
      vP := [seq(P||i, i = 1 ..nargs)];
      add(v[2]·mul(kappa[seq(degree(x, vP[i]), i = 1 ..nops(vP))], x = v[1]), v = makeTab(args))
  end proc:
```

Example: The joint moment $\mu_{2,1}$ in terms of joint cumulants is

```
> rtc(2, 1)
```

$$\kappa_{1,0}^{2}\,\kappa_{0,1} + \kappa_{2,0}\,\kappa_{0,1} + 2\,\kappa_{1,1}\,\kappa_{1,0} + \kappa_{2,1} \qquad \textbf{(4.1.2)}$$

## Product of Augmented Symmetric Functions

Products of multivariate polykays return products of power sums. For example the expectation of $k_2\ k_1$ is:

$$E[k_2\ k_1] = E\left[\left(\frac{n\,S_2 - S_1^2}{n\,(n-1)}\right)\cdot\left(\frac{S_1}{n}\right)\right] = \frac{E\left[S_2 S_1\right] n - E\left[S_1^3\right]}{n^2\,(n-1)}$$

where $E$ denotes the expectation. To express $E\left[S_2 S_1\right]$ and $E\left[S_1^3\right]$ in terms of moments, a fundamental result of estimation might be used if these products are transformed in augmented symmetric functions, see [9]. For univariate sample, an augmented symmetric function involves products of variables with different indexes, that is

$$A_{[i,j,k,\ldots]} = \sum_{a \neq b \neq c \neq \ldots} X_a^i X_b^j X_c^k \ldots$$

The *pSS* function carries out this task.

```
> uRv := proc(u, v)
  local w, i;
  w := [sort([op(u[1]), v·'X']), u[2]],
    seq(`if`(has(u[1,i],'X'), NULL, [sort([op(u[1,1..(i-1)]), u[1,i] + v·'X', op(u[1,(i+1)..nops(u[1])])])],
      u[2]]), i = 1..nops(u[1]));
    w;
  end proc:

  combV := proc(pU, V)
      local U, v, u;
      U := [[pU, 1]];
      for v in V do
          U := [seq(uRv(u, v), u = U)]
      end do;
      add(u[2]·'A'[sort(u[1])], u = eval(U, ['X' = 1]))
  end proc:

  pS := proc(uS, vS)
      local u, v;
      u := coeffs(uS); v := coeffs(vS);
      u·v·combV([op(op(uS/u))], [op(op(vS/v))])
  end proc:

  pSS := proc()
      local u, v, U, V;
      V := sort([args], (proc(x, y) evalb(nops(x) > nops(y)) end));
      U := pS(S[V[1]], S[V[2]]);
      for v in V[3..nops(V)] do U := add(pS(u, S[v]), u = [op(U)]) od;
      U;
  end proc:
```

Example: product of power sums $S_1 \cdot S_2^3$ in terms of augmented symmetric functions:

```
> pSS([1], [2], [2], [2])
```

$$A_{[1,2,2,2]} + 3\,A_{[2,2,3]} + 3\,A_{[1,2,4]} + 3\,A_{[3,4]} + 3\,A_{[2,5]} + A_{[1,6]} + A_{[7]} \qquad \textbf{(4.2.1)}$$

Example: product of power sums $S_{2,1} \cdot S_{1,1} \cdot S_{2,3}$ in terms of multivariate augmented symmetric functions:

```
> pSS([[2, 1]], [[1, 1]], [[2, 3]])
```

$$A_{[[1,1],[2,1],[2,3]]} + A_{[[2,3],[3,2]]} + A_{[[1,1],[4,4]]} + A_{[[2,1],[3,4]]} + A_{[[5,5]]} \qquad \textbf{(4.2.2)}$$

The following is a test explaing how to transform the power sum product $S_{1,1} \cdot S_{1,2}$ with in augmented symmetric functions with $n=3$:

```
> F1 := (sum(x[i]*y[i], i = 1..3))*(sum(x[i]*y[i]^2, i = 1..3)) : F1 := sort(expand(evalf(F1)))
```

$$F1 := x_1^2 y_1^3 + x_1 x_2 y_1^2 y_2 + x_1 x_2 y_1 y_2^2 + x_1 x_3 y_1^2 y_3 + x_1 x_3 y_1 y_3^2 + x_2^2 y_2^3 + x_2 x_3 y_2^2 y_3 + x_2 x_3 y_2 y_3^2 + x_3^2 y_3^3 \quad (4.2.3)$$

> $pSS([[1,1]],[[1,2]])$

$$A_{[[1,1],[1,2]]} + A_{[[2,3]]} \quad (4.2.4)$$

> $F2 := \sum_{j=1}^{3}\left(\sum_{i=1}^{3} x_i y_i x_j y_j^2\right) + \sum_{i=1}^{3} x_i^2 y_i^3 :$

$F2 := sort\left(add\left(add\left(`if`\left(i=j, 0, x_i y_i x_j y_j^2\right), i=1..3\right), j=1..3\right) + add\left(x_i^2 y_i^3, i=1..3\right)\right)$

$$F2 := x_1^2 y_1^3 + x_1 x_2 y_1^2 y_2 + x_1 x_2 y_1 y_2^2 + x_1 x_3 y_1^2 y_3 + x_1 x_3 y_1 y_3^2 + x_2^2 y_2^3 + x_2 x_3 y_2^2 y_3 + x_2 x_3 y_2 y_3^2 + x_3^2 y_3^3 \quad (4.2.5)$$

> $evalb(F2 = F1);$

$$true \quad (4.2.6)$$

## Expected values of Augmented Simmetric Functions

The fundamental result of estimation we refer to allows to express expectation of augmented symmetric functions in terms of moments [10], that is

$$E\left[A_{[i,j,k,\ldots]}\right] = n(n\text{-}1)\ldots(n\text{-}s\text{+}1)\mu_i\,\mu_j\,\mu_k\ldots$$

where $s$ denotes the number of elements in the list $[i, j, k, \ldots]$. The product $n(n\text{-}1)\ldots(n\text{-}s\text{+}1)$ is the lower factorial $(n)_s$. The $df$ function computes the lower factorial. This result can be suitably generalized to the multivariate case.

> $df := \mathbf{proc}(m)$
> $\quad \mathbf{local}\ i;$
> $\quad mul((n-i), i=0..m-1)$
> $\mathbf{end}:$

Example: the decreasing factorial $(n)_3$

> $df(3)$

$$n\ (n-1)\ (n-2) \quad (4.3.1)$$

The $E$ function computes the expectation of augmented symmetric functions.

> $E := \mathbf{proc}()$
> $\quad \mathbf{local}\ v, vArgs, u, r;$
> $\quad \mathbf{if}\ nargs = 1\ \mathbf{then return}\ n \cdot \mu_{op(args)}\ \mathbf{end if};$
> $\quad vArgs := seq([u], u = [args]);$
> $\quad v := seq\left(\left[op(u)_1/2, op\left(op(u)_2\right)\right], u = [op(2 \cdot pSS(vArgs))]\right);$
> $\quad add\left(df\left(nops\left(u_2\right)\right) \cdot u_1 \cdot mul\left(\mu_{op(r)}, r = u_2\right), u = v\right)$
> $\mathbf{end\ proc}:$

Example: the expectation of $A_{[[1,1],[2,1]]} + A_{[[3,2]]}$ is

> $E([1,1],[2,1]) + E([3,2])$

$$n\,(n-1)\,\mu_{1,1}\,\mu_{2,1}+2\,n\,\mu_{3,2} \qquad \textbf{(4.3.2)}$$

The *fE* function repeatedly executes the *E* function on an expression.

```
> fE := proc(s)
      local sn, sd, vI, vE, u, w;
      sd := denom(s);
      sn := expand(numer(s));
      vI := indets(s) minus {n};
      vE := [seq(u = 1, u = vI)];
      if nops(vI) = 1 or type(sn, `·`)
          then sn := [seq([[seq([[op(w)], degree(u, w)], w = vI)], eval(u, vE)], u = [sn])]

          else sn := [seq([[seq([[op(w)], degree(u, w)], w = vI)], eval(u, vE)], u = sn)]
      end if;
      sn := [seq([[seq(`if`(w[2] = 0, NULL, w), w = u[1])], u[2]], u = sn)];
      expand(add(E(seq(w[1]$w[2], w = u[1]))·u[2], u = sn))/sd
   end proc:
>
```

Example: the expectation of $\left(\frac{n\,S_2-S_1^2}{n\,(n-1)}\right)\cdot\left(\frac{S_1}{n}\right)$ is

```
> fE(polyk([2])·polyk([1]));
```

$$\frac{-n^3\,\mu_1^3+n^3\,\mu_1\,\mu_2+3\,n^2\,\mu_1^3-4\,n^2\,\mu_1\,\mu_2-2\,n\,\mu_1^3+n^2\,\mu_3+3\,n\,\mu_1\,\mu_2-n\,\mu_3}{n^2\,(n-1)} \qquad \textbf{(4.3.3)}$$

Example: the expectation of $S_3^2\cdot S_2+S_4$ is

```
> fE(S[3]^2·S[2]+S[4])
```

$$n^3\,\mu_2\,\mu_3^2-3\,n^2\,\mu_2\,\mu_3^2+n^2\,\mu_2\,\mu_6+2\,n^2\,\mu_3\,\mu_5+2\,n\,\mu_2\,\mu_3^2-n\,\mu_2\,\mu_6-2\,n\,\mu_3\,\mu_5+n\,\mu_4+n\,\mu_8 \qquad \textbf{(4.3.4)}$$

# Joint Cumulant of polykays: the main function

The *jcks* function returns the joint cumulants of multivariate polykays according to the procedure described in the previous section.

```
> jcks := proc(vK, vC)
   local s, vE, nv, dv, vS, evK, i, u, v, typeStru, vKi;
      vKi := vK1;
      if (nops(vKi) = 1) then if nops(vKi[1]) = 1 then typeStru := "ks" else typeStru := "km"; fi;
                              elif add(`if`(nops(op(u)) = 1, 0, 1), u = vK) = 0 then typeStru := "ps"
      else typeStru := "pm"; fi;
      if typeStru = "ks" or typeStru = "km"
         then evK := [seq(polyk(op(vK[i])), i = 1..nops(vK))];
         else evK := [seq(polyk(vK[i]), i = 1..nops(vK))]; fi;
      s := ctr(op(vC));
```

```
  vE := [seq(u = fE(mul(evK_i^[op(u)], i = 1 ..nops([op(u)]))), u = indets(s))];
  s := simplify(eval(s, vE));
  vE := [seq(u = rtc(op(u)), u = [seq(`if`(op(0, u) = mu, u, NULL), u = indets(s))])];
  s := simplify(eval(s, vE));
  nv := expand(numer(s));
  dv := denom(s);
  vS := [seq(v = 1, v = indets(s) minus {n})];
  if nops(vS) = 1 then return s end if;
  vS := [seq([u1/u2, u2], u = [seq([v, eval(v, vS)], v = nv)])];
  vS := {seq([v1, add(`if`(u1 = v1, u2, 0), u = vS)], v = vS)};
  add(v1*simplify(v2/dv), v = vS)
end proc:
```

## Examples

$\kappa_1(k_2)$ is the first cumulant of $k_2$, that can be computed as

```
> jcks([[[2]]], [1])
```

$$\kappa_2 \tag{5.1.1}$$

$\kappa_2(k_2)$ is the second cumulant of $k_2$, that can be computed as

```
> jcks([[[2]]], [2])
```

$$\frac{\kappa_4}{n} + \frac{2\,\kappa_2^2}{n-1} \tag{5.1.2}$$

$\kappa_3(k_2)$ is the third cumulant of $k_2$, that can be computed as

```
> jcks([[[2]]], [3])
```

$$\frac{\kappa_6}{n^2} + \frac{8\,\kappa_2^3}{(n-1)^2} + \frac{\kappa_3^2\,(4n-8)}{n\,(n-1)^2} + \frac{12\,\kappa_2\,\kappa_4}{n\,(n-1)} \tag{5.1.3}$$

In the following the variances of the first few *k*-statistics are computed, see [11]. Note that Var(X) = Cov(X, X).

```
> var(k1) = jcks([[[1]], [[1]]], [1, 1])
```

$$var(k_1) = \frac{\kappa_2}{n} \tag{5.1.4}$$

```
> var(k2) = jcks([[[2]], [[2]]], [1, 1])
```

$$var(k_2) = \frac{\kappa_4}{n} + \frac{2\,\kappa_2^2}{n-1} \tag{5.1.5}$$

```
> var(k3) = jcks([[[3]], [[3]]], [1, 1])
```

$$var(k_3) = \frac{\kappa_6}{n} + \frac{6\,\kappa_2^3\,n}{(n-1)\,(n-2)} + \frac{9\,\kappa_3^2}{n-1} + \frac{9\,\kappa_2\,\kappa_4}{n-1} \tag{5.1.6}$$

> $var(k_4) = jcks([\ [[4]], [[4]]\ ], [1, 1])$

$$var(k_4) = \frac{\kappa_8}{n} + \frac{24\,\kappa_2^4\,n\,(1+n)}{(n-1)\,(n-2)\,(n-3)} + \frac{34\,\kappa_4^2}{n-1} + \frac{144\,\kappa_2\,\kappa_3^2\,n}{(n-1)\,(n-2)} + \frac{16\,\kappa_2\,\kappa_6}{n-1} + \frac{72\,\kappa_2^2\,\kappa_4\,n}{(n-1)\,(n-2)} + \frac{48\,\kappa_3\,\kappa_5}{n-1} \tag{5.1.7}$$

More examples are given in [10], pag. 265.
The joint cumulant $\kappa(2^2\,1)$ of order (2, 1) of $(k_2, k_1)$ is

> $jcks([\ [[2]], [[1]]\ ], [2, 1])$

$$\frac{\kappa_5}{n^2} + \frac{4\,\kappa_2\,\kappa_3}{n\,(n-1)} \tag{5.1.8}$$

The joint cumulant $\kappa(2^2\,1^2)$ of order (2, 2) of $(k_2, k_1)$ is

> $jcks([\ [[2]], [[1]]\ ], [2, 2])$

$$\frac{\kappa_6}{n^3} + \frac{4\,\kappa_3^2}{n^2\,(n-1)} + \frac{4\,\kappa_2\,\kappa_4}{n^2\,(n-1)} \tag{5.1.9}$$

The joint cumulant $\kappa(3^1\,2^1)$ of order (1, 1) of $(k_3, k_2)$ is (see[11] pag. 271)

> $jcks([\ [[3]], [[2]]\ ], [1, 1])$

$$\frac{\kappa_5}{n} + \frac{6\,\kappa_2\,\kappa_3}{n-1} \tag{5.1.10}$$

The joint cumulant $\kappa(4^1\,2^2)$ of order (1,2) of $(k_4, k_2)$ is (see[11] pag. 272)

> $jcks([\ [[4]], [[2]]\ ], [1, 2])$

$$\frac{\kappa_8}{n^2} + \frac{\kappa_4^2\,(28\,n-40)}{n\,(n-1)^2} + \frac{120\,\kappa_2\,\kappa_3^2}{(n-1)^2} + \frac{20\,\kappa_2\,\kappa_6}{n\,(n-1)} + \frac{80\,\kappa_2^2\,\kappa_4}{(n-1)^2} + \frac{\kappa_3\,\kappa_5\,(40\,n-56)}{n\,(n-1)^2} \tag{5.1.11}$$

The joint cumulant of order (1,1) of $(k^{r,s}, k^{t,u})$ is (see[8] pag.94)

> $jcks([\ [[1, 1, 0, 0]], [[0, 0, 1, 1]]\ ], [1, 1])$

$$\frac{\kappa_{1,1,1,1}}{n} + \frac{\kappa_{1,0,0,1}\,\kappa_{0,1,1,0}}{n-1} + \frac{\kappa_{1,0,1,0}\,\kappa_{0,1,0,1}}{n-1} \tag{5.1.12}$$

The joint cumulant of order (1,1,1) of $\kappa_k(r, s|t, u|v, w)$ is (see[8] pag.94)

> *jcks*([ [[1, 1, 0, 0, 0, 0]], [[0, 0, 1, 1, 0, 0]], [[0, 0, 0, 0, 1, 1]] ], [1, 1, 1])

$$\frac{\kappa_{1,0,0,0,0,1}\,\kappa_{0,1,1,1,1,0}}{n\,(n-1)}+\frac{\kappa_{1,0,0,0,1,0}\,\kappa_{0,1,1,1,0,1}}{n\,(n-1)}+\frac{\kappa_{1,0,0,1,0,0}\,\kappa_{0,1,1,0,1,1}}{n\,(n-1)}$$
$$+\frac{\kappa_{1,0,0,1,1,1}\,\kappa_{0,1,1,0,0,0}}{n\,(n-1)}+\frac{\kappa_{1,0,1,0,0,0}\,\kappa_{0,1,0,1,1,1}}{n\,(n-1)}+\frac{\kappa_{1,0,1,0,1,1}\,\kappa_{0,1,0,1,0,0}}{n\,(n-1)}$$
$$+\frac{\kappa_{1,0,1,1,0,1}\,\kappa_{0,1,0,0,1,0}}{n\,(n-1)}+\frac{\kappa_{1,0,1,1,1,0}\,\kappa_{0,1,0,0,0,1}}{n\,(n-1)}+\frac{\kappa_{1,1,0,1,0,1}\,\kappa_{0,0,1,0,1,0}}{n\,(n-1)}$$
$$+\frac{\kappa_{1,1,0,1,1,0}\,\kappa_{0,0,1,0,0,1}}{n\,(n-1)}+\frac{\kappa_{1,1,1,0,0,1}\,\kappa_{0,0,0,1,1,0}}{n\,(n-1)}+\frac{\kappa_{1,1,1,0,1,0}\,\kappa_{0,0,0,1,0,1}}{n\,(n-1)}$$
$$+\frac{\kappa_{1,0,0,0,0,1}\,\kappa_{0,1,0,1,0,0}\,\kappa_{0,0,1,0,1,0}}{(n-1)^2}+\frac{\kappa_{1,0,0,0,0,1}\,\kappa_{0,1,1,0,0,0}\,\kappa_{0,0,0,1,1,0}}{(n-1)^2}$$
$$+\frac{\kappa_{1,0,0,0,1,0}\,\kappa_{0,1,0,1,0,0}\,\kappa_{0,0,1,0,0,1}}{(n-1)^2}+\frac{\kappa_{1,0,0,0,1,0}\,\kappa_{0,1,1,0,0,0}\,\kappa_{0,0,0,1,0,1}}{(n-1)^2}$$
$$+\frac{\kappa_{1,0,0,1,0,0}\,\kappa_{0,1,0,0,0,1}\,\kappa_{0,0,1,0,1,0}}{(n-1)^2}+\frac{\kappa_{1,0,0,1,0,0}\,\kappa_{0,1,0,0,1,0}\,\kappa_{0,0,1,0,0,1}}{(n-1)^2}$$
$$+\frac{\kappa_{1,0,1,0,0,0}\,\kappa_{0,1,0,0,0,1}\,\kappa_{0,0,0,1,1,0}}{(n-1)^2}+\frac{\kappa_{1,0,1,0,0,0}\,\kappa_{0,1,0,0,1,0}\,\kappa_{0,0,0,1,0,1}}{(n-1)^2}$$
$$+\frac{\kappa_{1,0,0,1,0,1}\,\kappa_{0,1,1,0,1,0}\,(n-2)}{n\,(n-1)^2}+\frac{\kappa_{1,0,0,1,1,0}\,\kappa_{0,1,1,0,0,1}\,(n-2)}{n\,(n-1)^2}$$
$$+\frac{\kappa_{1,0,1,0,0,1}\,\kappa_{0,1,0,1,1,0}\,(n-2)}{n\,(n-1)^2}+\frac{\kappa_{1,0,1,0,1,0}\,\kappa_{0,1,0,1,0,1}\,(n-2)}{n\,(n-1)^2}+\frac{\kappa_{1,1,1,1,1,1}}{n^2}$$

(5.1.13)

$\kappa_2(k_3)$ is the second cumulant of $k_3$ (the $k-$*statistic* of order 3 ), that can be computed as

> *jcks*([[[3]]], [2])

$$\frac{\kappa_6}{n}+\frac{6\,\kappa_2^3\,n}{(n-1)\,(n-2)}+\frac{9\,\kappa_3^2}{n-1}+\frac{9\,\kappa_2\,\kappa_4}{n-1}$$

(5.1.14)

$\kappa_2(k_{2,1})$ is the second cumulant of
$k_{2,1}$ (the multivariate $k-$statistics of order $(2, 1)$), that can be computed as

> *jcks*([[[2, 1]]], [2])

$$\frac{\kappa_{4,2}}{n}+\frac{5\,\kappa_{2,1}^2}{n-1}+\frac{4\,\kappa_{1,1}\,\kappa_{3,1}}{n-1}+\frac{4\,\kappa_{1,1}^2\,\kappa_{2,0}\,n}{(n-1)\,(n-2)}+\frac{4\,\kappa_{1,2}\,\kappa_{3,0}}{n-1}+\frac{2\,\kappa_{2,0}^2\,\kappa_{0,2}\,n}{(n-1)\,(n-2)}$$

(5.1.15)

$$+\frac{4\,\kappa_{2,2}\,\kappa_{2,0}}{n-1}+\frac{\kappa_{4,0}\,\kappa_{0,2}}{n-1}$$

$\kappa_2\left(k_{2;1}\right)$ is the second cumulant of $k_{2;1}$ (the polykay of order $(2,1)$), that can be computed as

> *jcks*([[[2], [1]]], [2])

$$\frac{\kappa_2^3\,(n^2-n+4)}{n\,(n-1)\,(n-2)}+\frac{\kappa_3^2}{n\,(n-1)}+\frac{2\,\kappa_1^2\,\kappa_2^2}{n-1}+\frac{\kappa_1^2\,\kappa_4}{n}+\frac{\kappa_2\,\kappa_4}{n\,(n-1)}+\frac{\kappa_1\,\kappa_2\,\kappa_3\,(2n-6)}{n\,(n-1)} \quad \textbf{(5.1.16)}$$

$\kappa_2\left(k_{1,1;1,0}\right)$ is the second cumulant of
$k_{1,1;1,0}$ (multivariate polykay of order $([1,1],[1,0])$), that can be computed as

> *jcks*([[[1, 1], [1, 0]]], [2])

$$\frac{\kappa_{2,1}^2}{n\,(n-1)}+\frac{\kappa_{1,0}^2\,\kappa_{1,1}^2}{n-1}+\frac{\kappa_{1,0}^2\,\kappa_{2,2}}{n}+\frac{\kappa_{1,1}^2\,\kappa_{2,0}\,(n^2-2n+4)}{n\,(n-1)\,(n-2)}+\frac{\kappa_{2,0}\,\kappa_{2,2}}{n\,(n-1)} \quad \textbf{(5.1.17)}$$

$$+\frac{\kappa_{2,0}^2\,\kappa_{0,2}}{(n-1)\,(n-2)}+\frac{\kappa_{0,2}\,\kappa_{1,0}^2\,\kappa_{2,0}}{n-1}+\frac{\kappa_{1,0}\,\kappa_{1,1}\,\kappa_{2,1}\,(2n-4)}{(n-1)\,n}-\frac{2\,\kappa_{1,0}\,\kappa_{1,2}\,\kappa_{2,0}}{n\,(n-1)}$$

The joint cumulant of order (1,1) of *simple polykays* $\left(k_{1;1},k_{1;2}\right)$ ]

> *jcks*([[[1], [1]], [[1], [2]]], [1, 1]);

$$\frac{\kappa_1\,\kappa_2^2\,(2n-4)}{(n-1)\,n}+\frac{2\,\kappa_1^2\,\kappa_3}{n}+\frac{2\,\kappa_2\,\kappa_3}{n\,(n-1)} \quad \textbf{(5.1.18)}$$

The joint cumulant of order (1,1) of $\left(k_{1,1;1,0},k_{2,0;0,1}\right)$

> *jcks*([[[1, 1], [1, 0]], [[2, 0], [0, 1]]], [1, 1])

$$\frac{\kappa_{1,1}\,\kappa_{3,1}}{n\,(n-1)}+\frac{\kappa_{1,1}^2\,\kappa_{2,0}\,(n^2-n+2)}{n\,(n-1)\,(n-2)}+\frac{\kappa_{1,2}\,\kappa_{3,0}}{n\,(n-1)}+\frac{2\,\kappa_{2,0}^2\,\kappa_{0,2}}{n\,(n-1)\,(n-2)}+\frac{\kappa_{0,1}\,\kappa_{1,0}\,\kappa_{3,1}}{n} \quad \textbf{(5.1.19)}$$

$$+\frac{\kappa_{0,1}\,\kappa_{1,1}\,\kappa_{3,0}}{n}-\frac{2\,\kappa_{0,1}\,\kappa_{2,0}\,\kappa_{2,1}}{n\,(n-1)}-\frac{\kappa_{0,2}\,\kappa_{1,0}\,\kappa_{3,0}}{n\,(n-1)}-\frac{\kappa_{1,0}\,\kappa_{1,1}\,\kappa_{2,1}}{n\,(n-1)}+\frac{\kappa_{1,0}\,\kappa_{1,2}\,\kappa_{2,0}}{n}$$

$$+\frac{2\,\kappa_{1,1}\,\kappa_{0,1}\,\kappa_{1,0}\,\kappa_{2,0}}{n-1}$$

If the variable *echo_type* is set to *true*, a message is displayed during each execution indicating which type of statistic is being computed.

*[ks = simple kStatistics, km = multivariate kStatistics, ps = simple polykays, pm = multivariate polykays]*

> *echo_type_KS* := *true*;

$$\textit{echo_type_KS} := \textit{true} \quad \textbf{(5.1.20)}$$

We will revisit some examples seen earlier.

$\kappa_2(k_2)$ is the second cumulant of $k_2$, that can be computed as

> *jcks*([ [[2]] ], [2])

$$execute_ks(2)$$

$$\frac{\kappa_4}{n} + \frac{2\,\kappa_2^2}{n-1} \tag{5.1.21}$$

$\kappa_2(k_{2,1})$ is the second cumulant of
$k_{2,1}$ (the multivariate $k-$statistics of order $(2, 1)$), that can be computed as

> *jcks*([[[2, 1]]], [2])

$$execute_km(2, 1)$$

$$\frac{\kappa_{4,2}}{n} + \frac{5\,\kappa_{2,1}^2}{n-1} + \frac{4\,\kappa_{1,1}\,\kappa_{3,1}}{n-1} + \frac{4\,\kappa_{1,1}^2\,\kappa_{2,0}\,n}{(n-1)\,(n-2)} + \frac{4\,\kappa_{1,2}\,\kappa_{3,0}}{n-1} + \frac{2\,\kappa_{2,0}^2\,\kappa_{0,2}\,n}{(n-1)\,(n-2)} \tag{5.1.22}$$
$$+ \frac{4\,\kappa_{2,2}\,\kappa_{2,0}}{n-1} + \frac{\kappa_{4,0}\,\kappa_{0,2}}{n-1}$$

The joint cumulant of order (1,1) of *simple polykays* $(k_{1;1}, k_{1;2})$ ]

> *jcks*([[[1], [1]], [[1], [2]]], [1, 1])

$$execute_ps(1, 1)$$

$$execute_ps(1, 2)$$

$$\frac{\kappa_1\,\kappa_2^2\,(2n-4)}{(n-1)\,n} + \frac{2\,\kappa_1^2\,\kappa_3}{n} + \frac{2\,\kappa_2\,\kappa_3}{n\,(n-1)} \tag{5.1.23}$$

The joint cumulant of order (1,1) of $(k_{1,1;1,0}, k_{2,0;0,1})$

> *jcks*([[[1, 1], [1, 0]], [[2, 0], [0, 1]]], [1, 1])

$$execute_pm([1, 1], [1, 0])$$

$$execute_pm([2, 0], [0, 1])$$

$$\frac{\kappa_{1,1}\,\kappa_{3,1}}{n\,(n-1)} + \frac{\kappa_{1,1}^2\,\kappa_{2,0}\,(n^2-n+2)}{n\,(n-1)\,(n-2)} + \frac{\kappa_{1,2}\,\kappa_{3,0}}{n\,(n-1)} + \frac{2\,\kappa_{2,0}^2\,\kappa_{0,2}}{n\,(n-1)\,(n-2)} + \frac{\kappa_{0,1}\,\kappa_{1,0}\,\kappa_{3,1}}{n} \tag{5.1.24}$$
$$+ \frac{\kappa_{0,1}\,\kappa_{1,1}\,\kappa_{3,0}}{n} - \frac{2\,\kappa_{0,1}\,\kappa_{2,0}\,\kappa_{2,1}}{n\,(n-1)} - \frac{\kappa_{0,2}\,\kappa_{1,0}\,\kappa_{3,0}}{n\,(n-1)} - \frac{\kappa_{1,0}\,\kappa_{1,1}\,\kappa_{2,1}}{n\,(n-1)} + \frac{\kappa_{1,0}\,\kappa_{1,2}\,\kappa_{2,0}}{n}$$
$$+ \frac{2\,\kappa_{1,1}\,\kappa_{0,1}\,\kappa_{1,0}\,\kappa_{2,0}}{n-1}$$

I reset the variable to the *false* state.

> *echo_type_KS* := *false*;

$$echo_type_KS := false \tag{5.1.25}$$

## Estimation of joint cumulants of polykays

Given a multivariate random sample, the *njcks* function estimates the joint cumulants of multivariate polykays. Internally, it invokes the *npolyk* function. The first two input parameters are the same as those of the *njcks* function, while the data matrix (or data vector, if appropriate) must be supplied as the third parameter.

The procedure first expresses the joint cumulants of a multivariate polykay in terms of the joint cumulants of the underlying population. It then replaces each occurrence of these joint cumulants, and of their products, with the corresponding numerical multivariate polykays, thereby providing an unbiased estimate of the desired joint cumulants.

```
> njcks := proc(vK, vC, vD)
    local s, u, sn, sd, vI, vE, N, w;
      s := jcks(vK, vC);
      if nops(op(vK[1])) = 1 then N := nops(op(vD)); else N := nops(vD); end if;
      sd := denom(s);
      sn := expand(numer(s));
      vI := indets(s) minus {n};
      vE := [seq(u = 1, u = vI)];
      if nops(vI) = 1 or type(sn, `*`) then
          sn := [seq([[seq([[op(w)], degree(u, w)], w = vI)], eval(u, vE)], u = [sn])]
      else
          sn := [seq([[seq([[op(w)], degree(u, w)], w = vI)], eval(u, vE)], u = sn)]
      end if;
      sn := [seq([[seq(`if`(w[2] = 0, NULL, w), w = u[1])], u[2]], u = sn)];
      eval(expand(add(npolyk([seq(w[1]$w[2], w = u[1])], vD)·u[2], u = sn))/sd, ['n' = N]);
  end:
```

Example: suppose to have the following (univariate) numerical sample

```
> data1 := [[16.34, 10.76, 11.84, 13.55, 15.85, 18.20, 7.51, 10.22, 12.52, 14.68, 16.08, 19.43,
      8.12, 11.20, 12.95, 14.77, 16.83, 19.80, 8.55, 11.58, 12.10, 15.02, 16.83, 16.98, 19.92, 9.47,
      11.68, 13.41, 15.35, 19.11]];
```

$$data1 := [[16.34, 10.76, 11.84, 13.55, 15.85, 18.20, 7.51, 10.22, 12.52, 14.68, 16.08, 19.43, 8.12, 11.20, 12.95, 14.77, 16.83, 19.80, 8.55, 11.58, 12.10, 15.02, 16.83, 16.98, 19.92, 9.47, 11.68, 13.41, 15.35, 19.11]] \quad \textbf{(5.2.1)}$$

The estimation of the joint cumulant of order (2,2) of $(k_2, k_1)$ is

```
> njcks([ [[2]], [[1]] ], [2, 2], data1)
```

$$0.01919781147 \quad \textbf{(5.2.2)}$$

Example: suppose to have the following (bivariate) numerical sample

```
> data2 := [[5.31, 11.16], [3.26, 3.26], [2.35, 2.35], [8.32, 14.34], [13.48, 49.45], [6.25, 15.05],
      [7.01, 7.01], [8.52, 8.52], [0.45, 0.45], [12.08, 12.08], [19.39, 10.42]];
```

$$data2 := [[5.31, 11.16], [3.26, 3.26], [2.35, 2.35], [8.32, 14.34], [13.48, 49.45], [6.25, \quad \textbf{(5.2.3)}$$

15.05 ], [ 7.01, 7.01 ], [ 8.52, 8.52 ], [ 0.45, 0.45 ], [ 12.08, 12.08 ], [ 19.39, 10.42 ] ]

An estimation of the joint cumulant of order (1,1) of $k_{1,1}$ $k_{1,2}$ is

> *njcks*( [ [ [ 1, 1 ] ], [ [ 1, 2 ] ] ], [ 1, 1 ], *data2* )

18250.79475 **(5.2.4)**

# Conclusions

The *jcks* function computes the joint cumulants of multivariate polykays. These routines are particularly useful for assessing the quality of the estimates provided by these estimators. For example, the variances of the first- through fourth-order *k*-statistics provide information about the precision of the estimates of the mean, variance, skewness, and kurtosis, respectively. Additional examples are presented in the final section of this worksheet.

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